

Equitable coloring extends Chernoff-Hoeffding bounds

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Outline

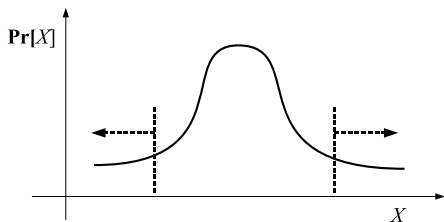
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- 2 A brief introduction to Chernoff-Hoeffding bounds
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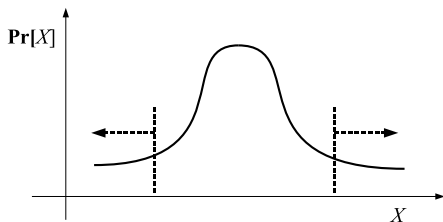
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- In 1952, Herman Chernoff introduced a technique which gives sharp upper bounds on the *tails* of the distribution of the **sum of mutually independent binary (Bernoulli) random variables**.
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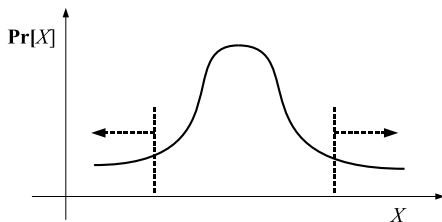
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Introduction (contd.)

- In many situations, tail probability bounds obtained using *Markov's inequality* or *Chebyshev's inequality* are too weak, while CH bounds are just right.
- CH bounds are extremely useful in design and analysis of randomized algorithms, in proofs by the probabilistic method, analysis in computational complexity, etc.
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Chernoff-Hoeffding bounds

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 - Assume that, for all i , $\Pr[X_i = 1] = p$ for some $p > 0$.
- We are interested in upper bounds on $\Pr[S \geq (1 + \delta)\mu]$ and $\Pr[S \leq (1 - \delta)\mu]$.
- Chernoff bounds lead to

$$\Pr[S \geq (1 + \delta)\mu] \leq \left(\frac{e^\delta}{(1 + \delta)^{(1 + \delta)}} \right)^\mu,$$

$$\Pr[S \leq (1 - \delta)\mu] \leq e^{-\mu\delta^2/2}.$$

When $\delta \leq 1$, we can derive $F^+(\mu, \delta) \leq e^{-\mu\delta^2/3}$.

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A simple application (a generous teacher and diligent students)

- There are n students, who work very hard all the time just like us. Their teacher, who is very generous, would like to reward them.
- In front of them, there is a sealed box which has 3 golden balls and 1 black ball inside.
- Each time one can pick a ball from the box and then put it back into the box (we assume that the students are honest).
- The teacher said he will treat the students a bountiful feast if more than $n/2$ students get golden balls.
- What is the probability that the students can't have a bountiful feast?

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A simple application (a generous teacher and diligent students) (contd.)

- For $i = 1, \dots, n$, $X_i = 1$: the i th student gets a golden ball; $X_i = 0$: the i th student gets a black ball.
- $\Pr[X_i = 1] = 3/4$ and $\Pr[X_i = 0] = 1/4$.
- Let $S = \sum_{i=1}^n X_i$. The event that the students have bad luck is $S \leq n/2$, and we have $\mu = \mathbf{E}[S] = 3n/4$.
- $\Pr[S \leq n/2] = \Pr[S \leq (1 - 1/3)\mu] \leq e^{-\mu(1/3)^2/2} = e^{-n/24}$.
- The probability is less than 0.66 if $n = 10$, less than 0.125 if $n = 50$, and less than 0.005 if $n = 130$.

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Hoeffding's extension

- Consider the case that X_i 's are mutually independent "bounded" random variables (i.e., $a_i \leq X_i \leq b_i$, for some positive real a_i and b_i).
- Hoeffding's extension of Chernoff's technique:

$$\Pr[|S - \mu| \geq \delta\mu] \leq 2e^{-2\mu^2\delta^2 / \sum_{i=1}^n (b_i - a_i)^2}.$$

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The crucial step and limitation of CH bounds

- A crucial step for deriving CH bounds is to calculate $\mathbf{E}[e^{tS}]$ for any positive real t (the moment generating function).

$$\mathbf{E}[e^{tS}] = \mathbf{E}[e^{t \sum_{i=1}^n X_i}] = \mathbf{E} \left[\prod_{i=1}^n e^{tX_i} \right] = \prod_{i=1}^n \mathbf{E}[e^{tX_i}].$$

- The last of the above equalities depends on the X_i 's being **mutually independent**.
- This is the limitation for CH bounds.
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Some basic definitions

- Let A be an event.
- A is said to be **mutually independent** of a set of events B_1, B_2, \dots, B_n if for any $I \subseteq \{1, 2, \dots, n\}$, $\Pr[A \mid \bigcap_{j \in I} B_j] = \Pr[A]$.

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Dependency graphs

- $\mathcal{X} = \{X_1, X_2, \dots, X_n\}$: a set of random variables.
- A **dependency graph** $G = (V, E)$ for \mathcal{X} has a vertex set $[n] = \{1, 2, \dots, n\}$ and for each i , X_i is mutually independent of the events $\{X_j \mid (i, j) \notin E\}$.
- We say that \mathcal{X} exhibits **d -bounded dependence** if \mathcal{X} has a dependency graph with maximum degree d .

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Note

- Let G be a dependency graph of \mathcal{X} .
- Assume that X_1, X_2, \dots, X_k correspond to an independent set of G .

$$\Pr[X_1 \mid X_2 \cap X_3 \cap \dots \cap X_k] = \frac{\Pr[X_1 \cap X_2 \cap X_3 \cap \dots \cap X_k]}{\Pr[X_2 \cap X_3 \cap \dots \cap X_k]} = \Pr[X_1].$$

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 - ▷ Must the dependency graph of S contain 0 edge?
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Another example for figuring out dependency graphs

- Consider an experiment of flipping a fair coin twice. Let \mathcal{X} be the set of the following events.
 - ▶ X_1 : the first flip is head;
 - ▶ X_2 : the second flip is tail;
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- The events can be shown to be pairwise independent for each two of them.
- If a graph with three vertices has *at most one edge*, it must NOT be a dependency graph of \mathcal{X} .
- ANY graph with three vertices and at least two edges is a dependency graph of \mathcal{X} .

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The main theorem

Theorem 1

For identically distributed Bernoulli random variables X_i with d -bounded dependence, for any $0 < \delta \leq 1$, we have the upper tail probability bound

$$\Pr[S \geq (1 + \delta)\mu] \leq \frac{4(d + 1)}{e} F^+(\mu, \delta)^{\frac{1}{d+1}} = \frac{4(d + 1)}{e} e^{-\mu\delta^2/3(d+1)}$$

and the lower tail probability bound

$$\Pr[S \leq (1 - \delta)\mu] \leq \frac{4(d + 1)}{e} F^-(\mu, \delta)^{\frac{1}{d+1}} = \frac{4(d + 1)}{e} e^{-\mu\delta^2/2(d+1)}$$

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An example: a randomized algorithm for Maximum Independent Set in a regular graph

- Given a k -regular n -vertex graph G . The following steps compute a large independent set in G .

Step 1: Delete each vertex from G independently with probability $1 - 1/k$.

Step 2: For each remaining edge, delete one of its endpoints.

- The vertices that remain after Step 2 form an independent set of G .

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 - ▶ Let $A = \sum_i A_i$ be a r.v.: the number of vertices remaining after Step 1.
- Let B_j be an indicator r.v. such that $B_j = 1$ if edge e_j is *not* deleted in Step 1.
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- It is easy to see that $\mathbf{E}[A] = n/k$ and $\mathbf{E}[B] = (1/k)^2 \cdot kn/2 = n/2k$.

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- It is easy to see that $\mathbf{E}[A] = n/k$ and $\mathbf{E}[B] = (1/k)^2 \cdot kn/2 = n/2k$.

An example: a randomized algorithm for Maximum Independent Set in a regular graph (contd.)

- The size of the independent set computed by the algorithm:
 $\geq A - B$.
- Hence the expected size of the solution produced by the algorithm is
 $\geq n/2k$.
 - ▶ A randomized $O(1)$ -factor approximation algorithm for Maximum Independent Set.

An example: a randomized algorithm for Maximum Independent Set in a regular graph (contd.)

- Actually we can show that $A - B$ is very close to $n/2k$ with high probability.
- It is clear that A_i 's are mutually independent, so CH bounds can be applied.
- However, B_i 's are NOT mutually independent.
 - ▶ B_i is mutually independent of B_j 's if edge j 's are not incident on any endpoints of edge i .
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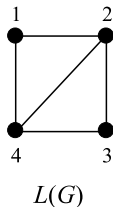
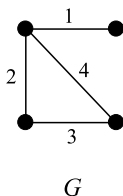
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An example: a randomized algorithm for Maximum Independent Set in a regular graph (contd.)

- The line graph (i.e., edge graph) $L(G)$ of G is a dependency graph of the B_i 's.
 - ▶ $L(G)$: every vertex of $L(G)$ represents an edge of G , and two vertices of $L(G)$ are adjacent iff their corresponding edge share a common endpoint in G .



An example: a randomized algorithm for Maximum Independent Set in a regular graph (contd.)

- G is k -regular $\longrightarrow L(G)$ is $2(k - 1)$ -regular $\longrightarrow B_i$'s exhibit $2(k - 1)$ -bounded dependence.
- $\mathbb{E}[B]/(2k - 1) = \Omega(n)$.
 - ▶ $\Omega(\log^{1+\rho} kn/2)$ for any $\rho > 0$.
- Thus the main theorem of this paper can be applied, and then we know the algorithm indeed produces a large independent set with high probability.

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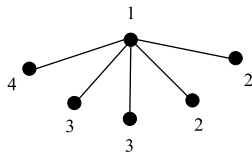
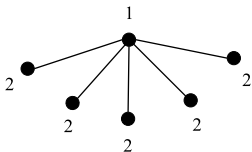
- 1 Introduction
- 2 A brief introduction to Chernoff-Hoeffding bounds
- 3 The main theorem and an illustrating example
- 4 Proof of the main theorem**
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t -equitable coloring

Definition 2

A coloring of a graph is **equitable** if the sizes of any pair of color classes are within one of each other.

- t -equitable coloring: an equitable coloring using t colors.



A deep result by Hajnal and Szemerédi

Hajnal-Szemerédi (1970)

A graph G with maximum degree Δ has a $(\Delta + 1)$ -equitable coloring.

Lemma 3

Suppose that X_i 's are identical Bernoulli random variables with dependency graph G , and suppose G has a t -equitable coloring. Then for any $0 < \delta \leq 1$, we have

$$\Pr[S \geq (1 + \delta)\mu] \leq \frac{4t}{e} F^+(\mu, \delta)^{1/t},$$
$$\Pr[S \leq (1 - \delta)\mu] \leq \frac{4t}{e} F^-(\mu, \delta)^{1/t}.$$

Theorem 4

Suppose the X_i 's are identical Bernoulli random variables exhibiting d -bounded dependence. Then, for any $0 < \delta \leq 1$, we have

$$\Pr[S \geq (1 + \delta)\mu] \leq \frac{4(d + 1)}{e} F^+(\mu, \delta)^{\frac{1}{d+1}},$$
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Proof of Lemma 3

- For convenience, assume that $\mathbf{E}[X_i] = \mu'$ for each i , and let $[t]$ denote $\{1, 2, \dots, t\}$.
- Let C_1, C_2, \dots, C_t be the t color classes in a t -equitable-coloring of G .
- For each $i \in [t]$, let $\mu_i = \mathbf{E}[\sum_{j \in C_i} X_j]$ (i.e., $\mu' |C_i|$).

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Proof of Lemma 3 (contd.)

$$\begin{aligned} S \geq (1 + \delta)\mu &\equiv S \geq (1 + \delta)\mu' n \\ &\equiv S \geq (1 + \delta)\mu' \sum_{i \in [t]} |C_i| \\ &\equiv S \geq \sum_{i \in [t]} (1 + \delta)\mu' |C_i| \\ &\equiv \sum_{i \in [t]} \sum_{j \in C_j} X_j \geq \sum_{i \in [t]} (1 + \delta)\mu_i. \end{aligned}$$

- The first equivalence: $\mu = \mathbf{E}[\sum_{i \in [n]} X_i] = \sum_{i \in [n]} \mathbf{E}[X_i] = n\mu'$.
- The second equivalence: C_i 's form a partition of $[n]$.
- The last equivalence: expressing S as the sum of the X_i 's grouped into color classes.

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- Thus the upper tail probability is proved.

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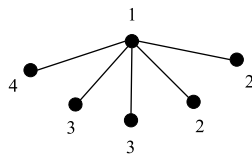
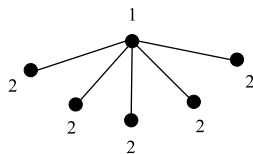
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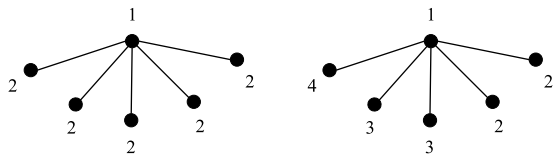
- $\chi(G)$: the chromatic number of G .
- $\chi_{eq}(G)$: the fewest colors required to equitably color the graph G .
- E.g., $\chi(G) = 2$ and $\chi_{eq}(G) = \lceil (n-1)/2 \rceil + 1$ when G is an n -vertex star graph.



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Bollobás-Guy (1983)

A tree T with n vertices is equitably 3-colorable if $n \geq 3\Delta(T) - 8$ or if $n = 3\Delta(T) - 10$.

- The theorem implies that if $\Delta(T) \leq n/3$, then T can be equitably 3-colored. Thus we have

Theorem 5

Suppose that X_i 's are identical Bernoulli random variables such that the corresponding dependency graph is a tree with maximum degree at most $n/3$. Then we have the following bounds

$$\Pr[S \geq (1 + \delta)\mu] \leq \frac{12}{e} F^+(\mu, \delta)^{1/3},$$
$$\Pr[S \leq (1 - \delta)\mu] \leq \frac{12}{e} F^-(\mu, \delta)^{1/3}.$$

Pemmaraju (2001); technical report

A connected outerplanar graph with n vertices and vertex degree at most $n/6$ has a 6-equitable coloring.

Theorem 6

Suppose that X_i 's are identical Bernoulli random variables whose dependency graph is outerplanar with maximum degree at most $n/6$. Then we have the following bounds

$$\Pr[S \geq (1 + \delta)\mu] \leq \frac{24}{e} F^+(\mu, \delta)^{1/6},$$

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Some further remarks

- Are the bounds on the vertex-degree required to obtain sharp bounds?
- a (c, α) -coloring: a vertex coloring such that
 - ▶ $\leq c$ vertices are not colored.
 - ▶ for any pair of color classes C and C' , $|C| \leq \alpha|C'|$.
- It is possible to extend Bollobás-Guy Theorem to have the following results.

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Every tree has a $(1, 5)$ -coloring with two colors.

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Thank you!