

# The $K$ -armed dueling bandits problem

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# Outline

- 1 The dueling bandits problem
- 2 The algorithm
- 3 The main analysis
  - Justification of the confidence intervals
  - Regret per match
  - Mistake bound
  - Exploration bound w.h.p.
  - Expected regret upper bound
- 4 The lower bound



# Motivations

- The conventional bandit problem :
  - Choose, in each of  $T$  iterations, one of the  $K$  possible bandits/arms/strategies  $\mathcal{B} = \{b_1, \dots, b_K\}$ .
  - Receive the payoff in  $[0, 1]$  (*initially unknown*) in each iteration.
  - **Goal:** Maximize the total payoff.
- It's difficult to elicit absolute-scale payoffs in some applications.
  - One can only rely on *relative* judgment of payoff.
- Given a collection of  $K$  bandits, we wish to find a sequence of noisy comparisons that has low regret.



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# Noisy comparisons

- $\Pr[b \succ b'] := \epsilon(b, b') + 1/2$ .
  - $\epsilon(b, b') \in (-1/2, 1/2)$ : a measure distinguishing  $b$  and  $b'$ .
    - $\epsilon(b, b') = -\epsilon(b', b)$
    - $\epsilon_{i,j} \equiv \epsilon(b_i, b_j)$ .
  - $b \succ b' \Rightarrow \epsilon(b, b') > 0$ .
- ★ The noisy comparisons are independent and  $\Pr[b \succ b']$  is stationary over time.



# Regrets

- $(b_1^{(t)}, b_2^{(t)})$ : the bandits chosen at iteration  $t$ .
- $b^*$ : the overall best bandit.
- $T$  be time horizon.

## Regrets

- The **strong regret**

$$R_T = \sum_{t=1}^T \max\{\epsilon(b^*, b_1^{(t)}), \epsilon(b^*, b_2^{(t)})\}.$$

- The **weak regret**

$$\tilde{R}_T = \sum_{t=1}^T \min\{\epsilon(b^*, b_1^{(t)}), \epsilon(b^*, b_2^{(t)})\}.$$





# Modeling assumptions

## Strong stochastic transitivity

For bandits  $b_i \succ b_j \succ b_k$ ,

$$\epsilon_{i,k} \geq \max\{\epsilon_{i,j}, \epsilon_{j,k}\}.$$

## Strong triangular inequality

For bandits  $b_i \succ b_j \succ b_k$ ,

$$\epsilon_{i,k} \leq \epsilon_{i,j} + \epsilon_{j,k}.$$



# The Algorithm



# Explore then exploit

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## Algorithm 1 Explore then exploit

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1: Input:  $T, \mathcal{B} = \{b_1, \dots, b_K\}$ , EXPLORE
2:  $(\hat{b}, \hat{T}) \leftarrow \text{EXPLORE}(T, \mathcal{B})$ 
3: for  $t = \hat{T} + 1, \dots, T$  do
4:   compare  $\hat{b}$  and  $\hat{b}$ 
5: end for

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## Algorithm 2 Interleaved Filter (IF).

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```

1: Input:  $T, \mathcal{B} = \{b_1, \dots, b_K\}$ 
2:  $\delta \leftarrow 1/(TK^2)$ 
3: Choose  $\hat{b} \in \mathcal{B}$  randomly
4:  $W \leftarrow \{b_1, \dots, b_K\} \setminus \{\hat{b}\}$ 
5:  $\forall b \in W$ , maintain estimate  $\hat{P}_{\hat{b},b}$  of  $P(\hat{b} > b)$  according to (6)
6:  $\forall b \in W$ , maintain  $1 - \delta$  confidence interval  $\hat{C}_{\hat{b},b}$  of  $\hat{P}_{\hat{b},b}$  according to (7), (8)
7: while  $W \neq \emptyset$  do
8:   for  $b \in W$  do
9:     compare  $\hat{b}$  and  $b$ 
10:    update  $\hat{P}_{\hat{b},b}, \hat{C}_{\hat{b},b}$ 
11:   end for
12:   while  $\exists b \in W$  s.t.  $(\hat{P}_{\hat{b},b} > 1/2 \wedge 1/2 \notin \hat{C}_{\hat{b},b})$  do
13:      $W \leftarrow W \setminus \{b\}$  //  $\hat{b}$  declared winner against  $b$ 
14:   end while
15:   if  $\exists b' \in W$  s.t.  $(\hat{P}_{\hat{b},b'} < 1/2 \wedge 1/2 \notin \hat{C}_{\hat{b},b'})$  then
16:     while  $\exists b \in W$  s.t.  $\hat{P}_{\hat{b},b} > 1/2$  do
17:        $W \leftarrow W \setminus \{b\}$  // pruning
18:     end while
19:      $\hat{b} \leftarrow b', W \leftarrow W \setminus \{b'\}$  //  $b'$  declared winner against  $\hat{b}$  (new round)
20:      $\forall b \in W$ , reset  $\hat{P}_{\hat{b},b}$  and  $\hat{C}_{\hat{b},b}$ 
21:   end if
22: end while
23:  $\hat{T} \leftarrow$  Total Comparisons Made
24: return  $(\hat{b}, \hat{T})$ 

```

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# The exploit algorithm

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22: end while
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24: return  $(\hat{b}, \hat{T})$ 

```

- $\hat{P}_{i,j} = \frac{\# b_i \text{ wins}}{\# \text{ comparisons}}$

The empirical estimate of  $\Pr[b_i \succ b_j]$  after  $t$  comparisons.

- Confidence interval:

$$\hat{C}_{i,j} := (\hat{P}_{i,j} - c_t, \hat{P}_{i,j} + c_t),$$

where  $c_t = \sqrt{4 \log(1/\delta)/t}$ .



# Contribution of this paper

## Theorem 1

Running Algorithm 1 with  $\mathcal{B} = \{b_1, \dots, b_K\}$ , time horizon  $T$  ( $T \geq K$ ), then **IF** incurs expected regret (weak & strong) bounded by

$$\mathbf{E}[R_T] = O(\mathbf{E}[R_T^{IF}]) = O\left(\frac{K}{\epsilon_{1,2}} \log T\right).$$

## Theorem 2

For any fixed  $\epsilon > 0$  and any algorithm  $\phi$  for the  $K$ -armed dueling bandit problem, there exists a problem instance such that

$$R_T^\phi = \Omega\left(\frac{K}{\epsilon} \log T\right),$$

where  $\epsilon = \min_{b \neq b^*} \Pr[b^* \succ b]$ .

# Crucial lemmas

## Lemma 1

The probability that **IF** makes a mistake resulting in the elimination of the best bandit  $b_1$  is  $\leq 1/T$ .

- $\mathbf{E}[R_T] \leq (1 - 1/T)\mathbf{E}[R_T^{IF}] + (1/T) \cdot O(T) = O(\mathbf{E}[R_T^{IF}])$ .
  - $R_T^{IF}$ : the regret incurred from **IF**.



## Crucial lemmas (contd.)

### Lemma 2

Assuming **IF** is mistake-free, then with high probability,

$$R_T^{IF} = O\left(\frac{K \log K}{\epsilon_{1,2}} \log T\right)$$

for both weak and strong regret.

### Lemma 3

Assuming **IF** is mistake-free, then

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## Some more terminologies

- **IF** makes a “mistake”: it draws a false conclusion regarding a bandit pair.
- A “match”: all the comparisons **IF** makes between two bandits.
- A “round”: all the matches played by the *incumbent* bandit  $\hat{b}$ .





# The Main Analysis



# Justification of the confidence intervals

## Lemma 4

- For  $\delta = 1/(TK^2)$ , the number of comparisons in a match b/w  $b_i, b_j$  is

$$O\left(\frac{1}{\epsilon_{i,j}^2} \log(TK)\right).$$

- $\Pr[\text{an inferior bandit is declared the winner at some time } t \leq T] \leq \delta.$

- IF makes a mistake at time  $t \Rightarrow 1/2 + \epsilon_{i,j} \notin \hat{C}_{i,j}.$

- Note:  $\mathbf{E}[\hat{P}_{i,j}] = 1/2 + \epsilon_{i,j}.$

- $\Pr[1/2 + \epsilon_{i,j} \notin \hat{C}_{i,j}] = \Pr[|\hat{P}_{i,j} - \mathbf{E}[\hat{P}_{i,j}]| \geq c_t] \leq 2 \cdot e^{-2t \cdot c_t^2} = 2/(T^8 K^{16}).$

- 

$$\Pr\left[\bigcup_{t=1}^T \{1/2 + \epsilon_{i,j} \notin \hat{C}_{i,j}\}\right] \leq \frac{2T}{T^8 K^{16}} \leq \frac{1}{TK^2} = \delta.$$



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- Note:  $\mathbf{E}[\hat{P}_{i,j}] = 1/2 + \epsilon_{i,j}.$

- $\Pr[1/2 + \epsilon_{i,j} \notin \hat{C}_{i,j}] = \Pr[|\hat{P}_{i,j} - \mathbf{E}[\hat{P}_{i,j}]| \geq c_t] \leq 2 \cdot e^{-2t \cdot c_t^2} = 2/(T^8 K^{16}).$

- $$\Pr\left[\bigcup_{t=1}^T \{1/2 + \epsilon_{i,j} \notin \hat{C}_{i,j}\}\right] \leq \frac{2T}{T^8 K^{16}} \leq \frac{1}{TK^2} = \delta.$$



# Proof of Lemma 4 (contd.)

- By the stopping condition of **IF**, the match terminates at any time  $t$  if  $\hat{P}_{i,j} - c_t > 1/2$ .
  - If  $n > t$ , then  $\hat{P}_{i,j} - c_t \leq 1/2$ .
- $\Pr[n > t] \leq \Pr[\hat{P}_{i,j} - c_t \leq 1/2] = \Pr[\hat{P}_{i,j} - 1/2 - \epsilon_{i,j} \leq c_t - \epsilon_{i,j}] = \Pr[\mathbf{E}[\hat{P}_{i,j}] - \hat{P}_{i,j} \geq \epsilon_{i,j} - c_t]$ .
- Set  $m \geq 8$  and  $t \geq \lceil 2m \log(TK^2)/\epsilon_{i,j}^2 \rceil$  (then  $c_t \leq \epsilon_{i,j}/2$ ), we will have

$$\Pr\left(n \geq \frac{m}{\epsilon_{i,j}^2} \log(TK)\right) \leq \frac{1}{(TK)^m}.$$



# Regret per match

## Lemma 5

Assume that  $b_1$  has not been removed and  $T \geq K$ , then w.h.p. the accumulated weak/strong regret **from any match** is

$$O\left(\frac{1}{\epsilon_{1,2}} \log T\right).$$

- Suppose  $\hat{b} = b_j$  is playing a match against  $b_i$ .
- By Lemma 4, any match played by  $b_j$  contains at most

$$O\left(\frac{1}{\epsilon_{1,j}^2} \log(TK)\right) = O\left(\frac{1}{\epsilon_{1,2}^2} \log(TK)\right) \text{ comparisons.}$$

- **Note:** All matches within a round are played *simultaneously*.



# Proof of Lemma 5 (contd.)

- The accumulated weak regret is bounded by

$$\begin{aligned}\epsilon_{1,j} \cdot O\left(\frac{1}{\epsilon_{1,j}^2} \log(TK)\right) &= O\left(\frac{1}{\epsilon_{1,j}} \log(TK)\right) \\ &= O\left(\frac{1}{\epsilon_{1,2}} \log(T)\right).\end{aligned}$$



# Mistake bound

- **IF** eliminates the best bandit  $b_1$  if
  - an inferior bandit defeats  $b_1$ , or
  - $b_1$  is removed during the pruning step (lines 16–18).
- Consider the second case.

## Lemma 6

For all triples of bandits  $b, b', \hat{b}$  such that  $b \succ b'$ , the probability that **IF** eliminates  $b$  in a pruning step, where

- $b'$  wins a match against  $\hat{b}$  while
- $b$  is empirically inferior to  $\hat{b}$ ,

is  $\leq \delta$ .



# Proof of Lemma 6

- $X_1, X_2, \dots$  : an infinite sequence of i.i.d. Bernoulli random variables with  $\mathbf{E}[X_i] = \Pr[\hat{b} \succ b']$ .
- $Y_1, Y_2, \dots$  : an infinite sequence of i.i.d. Bernoulli random variables with  $\mathbf{E}[Y_i] = \Pr[\hat{b} \succ b]$ .
  - $X_i$  (resp.  $Y_i$ ) represents the outcome of the  $i$ th comparison b/w  $\hat{b}$  &  $b'$  (resp.  $\hat{b}$  &  $b$ ).
- If  $b$  is eliminated in a pruning step at the end of a match consisting of  $n$  comparisons b/w  $b'$  and  $\hat{b}$ , then

$$X_1 + \dots + X_n < n/2 - \sqrt{4n \log(1/\delta)},$$

$$Y_1 + \dots + Y_n > n/2.$$





# Proof of Lemma 6 (contd.)

- Define  $Z_i = Y_i - X_i$ , we have

$$Z_1 + \dots + Z_n > \sqrt{4n \log(1/\delta)}.$$

- $(Z_i)_{i=1}^{\infty}$  are i.i.d., and  $|Z_i| \leq 1, \forall i$ .
- ★  $\mathbf{E}[Z_i] = \Pr[\hat{b} \succ b] - \Pr[\hat{b} \succ b'] \leq 0$ .
- Taking Hoeffding's inequality & union bound...



# Proof of Lemma 1

## Lemma 1

The probability that **IF** makes a mistake resulting in the elimination of the best bandit  $b_1$  is  $\leq 1/T$ .

- For every  $i$ , the probability that  $b_1$  is eliminated in a match against  $b_i$  is  $\leq \delta$  (Lemma 4).
- For all  $i, j$ , the probability that  $b_1$  is eliminated in a pruning step resulting from a match where  $b_i$  defeats  $b_j$  is  $\leq \delta$  (Lemma 6).
- ★ The probability that **IF** makes a mistake resulting in eliminating  $b_1$  is  $\leq \delta(K-1) + \delta(K-1)^2 < \delta K^2 = 1/T$ .



# The regret upper bound

## Lemma 2

Assuming **IF** is mistake-free, then with high probability,

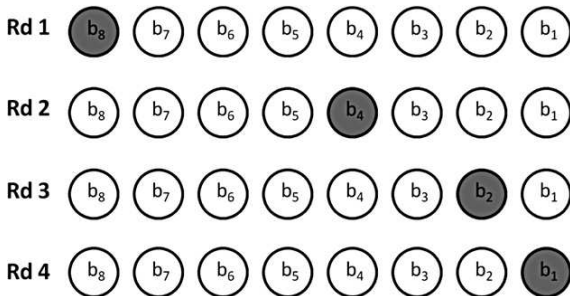
$$R_T^{IF} = O\left(\frac{K \log K}{\epsilon_{1,2}} \log T\right)$$

for both weak and strong regret.

- We wish to prove that the number of candidate bandits (i.e., # rounds) is  $O(\log K)$  w.h.p.
- Model the sequence of candidate bandits as a **random walk**.



# Random walk model



- $p_i$ : the prob.  $b_i$  will be the incumbent in the following round.
  - $p_{j-1} \leq \dots \leq p_1$  ( $\because$  strong stochastic transitivity).
- The “worst case”:  $p_{j-1} = \dots = p_1 = 1/(j-1)$  (assuming no mistakes are made).



# Random Walk Model (contd.)

## Random Walk Model

Define a random walk graph with  $K$  nodes labeled  $b_1, \dots, b_K$ . Each node  $b_j$  ( $j > 1$ ) transitions to  $b_i$  for  $j > i \geq 1$  with prob.  $1/(j-1)$  (uniform). The final node  $b_1$  is an absorbing node.

## Proposition 1

If  $S$  and  $\tilde{S}$  are random variables corresponding to the number of rounds in **IF** and the Random Walk Model, resp., then

$$\forall x : \Pr[S \geq x] \leq \Pr[\tilde{S} \geq x].$$



# Analysis of the Random Walk Model

## Lemma 7

Let  $X_i$  ( $1 \leq i \leq K$ ) be an indicator random variable corresponding to whether a random walk starting at  $b_K$  visits  $b_i$  in the Random Walk Model. Then

$$\Pr[X_i = 1] = \frac{1}{i},$$

and for all  $W \subseteq \{X_1, \dots, X_{K-1}\}$ ,

$$\Pr[\bigwedge_{i \in W} X_i] = \prod_{X_i \in W} \Pr[X_i].$$

- We can express the number of steps taken by a random walk from  $b_K$  to  $b_1$  as  $S_k = 1 + \sum_{i=1}^{K-1} X_i$ . Then,

$$\mathbf{E}[S_k] = 1 + \sum_{i=1}^{K-1} \mathbf{E}[X_i] = 1 + H_{K-1} \approx \log K.$$



# Analysis of the Random Walk Model (contd.)

## Lemma 8

Assuming **IF** is mistake-free, then it runs for  $O(\log K)$  rounds w.h.p..

## Corollary 1

Assuming **IF** is mistake-free, then it plays  $O(K \log K)$  matches w.h.p.

- $O(\log T / \epsilon_{1,2})$  accumulated regret per match (Lemma 5).
- ▷ Lemma 2 (i.e.,  $R_T^{IF} = O((K \log K) \log T / \epsilon_{1,2})$ ) follows.



# Expected Regret Upper Bound



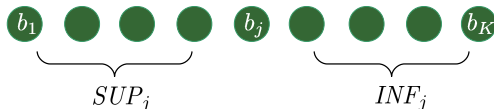


# Expected regret upper bound

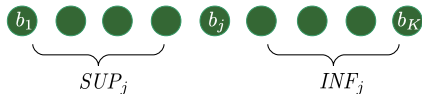
## Lemma 9

Assuming **IF** is mistake-free, then it plays  $O(K)$  matches *in expectation*.

- $B_j$ : # matches played by  $b_j$  when it is NOT the incumbent.
- $B_j = INF_j + SUP_j$ , where
  - $INF_j$ : # matches played by  $b_j$  against  $b_i$  for  $i > j$ .
  - $SUP_j$ : # matches played by  $b_j$  against  $b_i$  for  $i < j$ .
- Then  $\sum_{j=1}^K \mathbf{E}[B_j] = \sum_{j=1}^K (\mathbf{E}[INF_j] + \mathbf{E}[SUP_j])$ .



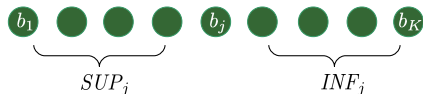
# Proof of Lemma 9 (contd.)



- $$\mathbf{E}[INF_j] \leq 1 + \sum_{i=j+1}^{K-1} \frac{1}{i} = 1 + H_{K-1} - H_j.$$



# Proof of Lemma 9 (contd.)



- Assume that  $b_j$  does NOT lose a match (not to be eliminated) to any superior incumbent  $b_i$  before  $b_i$  is defeated unless  $b_i = b_1$ .
- $\mathcal{E}_{j,t}$ :  $b_j$  is pruned after the  $t$ -th round where the incumbent bandit is superior to  $b_j$ , conditioned on NOT being pruned in the first  $t - 1$  such rounds.
- $G_{j,t}$ : # matches beyond the first  $t - 1$  played by  $b_j$  against a superior incumbent, conditioned on playing  $\geq t - 1$  such matches.
- $\mathbf{E}[G_{j,t}] = 1 + \Pr[\mathcal{E}_{j,t}^c] \cdot \mathbf{E}[G_{j,t+1}]$ .
- ★  $\mathbf{E}[SUP_j] \leq \mathbf{E}[G_{j,1}] = 1 + \Pr[\mathcal{E}_{j,1}^c] \cdot \mathbf{E}[G_{j,2}] \leq 1 + 1/2 + 1/4 + \dots = 2$ .  
 $(\because \Pr[\mathcal{E}_{j,t}] \leq 1/2, \forall j \neq 1, t)$



# Proof of Lemma 9 (contd.)

- Thus,

$$\begin{aligned}
 \sum_{j=1}^K (\mathbf{E}[INF_j] + \mathbf{E}[SUP_j]) &\leq \sum_{j=1}^K (1 + H_{K-1} - H_j) + 2K \\
 &= \sum_{j=1}^K \left( 1 + \sum_{i=j+1}^{K-1} \frac{1}{i} \right) + 2K \\
 &= \sum_{j=1}^K (j-1) \frac{1}{j} + 3K \\
 &= O(K).
 \end{aligned}$$



# The Lower Bound



# The lower bound

## Theorem 2

For any fixed  $\epsilon > 0$  and any algorithm  $\phi$  for the  $K$ -armed dueling bandit problem, there exists a problem instance such that

$$R_T^\phi = \Omega\left(\frac{K}{\epsilon} \log T\right),$$

where  $\epsilon = \min_{b \neq b^*} \Pr[b^* \succ b]$ .



# Construction of the problem instances

## A family of $K$ problem instances

- In instance  $j$ , let  $b_j$  be the best bandit, order the remaining ones by their indices.
  - In instance  $j$ , we have  $b_j \succ b_k$  for all  $k \neq j$  and we have  $b_i \succ b_k$  whenever  $i < k$ .
- $\Pr[b_i \succ b_k] := 1/2 + \epsilon$  whenever  $b_i \succ b_k$ .
- $q_j$ : the *distribution* on  $T$ -step histories induced by  $\phi$  under instance  $j$ .
  
- $n_{j,T}$ : # comparisons involving  $b_j$  scheduled by  $\phi$  up to time  $T$ .



# Proving the lower bound

## Lemma 10

Let  $\phi$  be an algorithm for the  $K$ -armed dueling bandits problem, such that  $R_T^\phi = o(T^a)$  for all  $a > 0$ . Then for all  $j$ ,

$$\mathbf{E}_{q_1}[n_{j,T}] = \Omega\left(\frac{\log T}{\epsilon^2}\right).$$

- If  $R_T^\phi \neq o(T^a)$ , then Theorem 2 holds trivially.
- On instance  $j$ ,  $\phi$  incurs regret  $\geq \epsilon$  every time when it plays a match involving  $b_j \neq b_1$ .

$$R_T^\phi \geq \sum_{j \neq 1} \epsilon \cdot \mathbf{E}_{q_1}[n_{j,T}] = \Omega\left(\frac{K}{\epsilon} \log T\right).$$





## Proof of Lemma 10

- $\mathcal{E}_j$ : the event that  $n_{j,T} < \log(T)/\epsilon^2$ .
- $J := \{j \mid q_1(\mathcal{E}_j) < 1/3\}$ .
- For each  $j \in J$ :

$$\mathbf{E}_{q_1}[n_{j,T}] \geq q_1(\mathcal{E}_j^c)(\log(T)/\epsilon^2) = \Omega\left(\frac{\log(T)}{\epsilon^2}\right).$$

- Hence, it remains to show that  $\mathbf{E}_{q_1}[n_{j,T}] = \Omega(\log(T)/\epsilon^2)$  for each  $j \notin J$ .



## Proof of Lemma 10

- $\mathcal{E}_j$ : the event that  $n_{j,T} < \log(T)/\epsilon^2$ .
- $J := \{j \mid q_1(\mathcal{E}_j) < 1/3\}$ .
- For each  $j \in J$ :

$$\mathbf{E}_{q_1}[n_{j,T}] \geq q_1(\mathcal{E}_j^c)(\log(T)/\epsilon^2) = \Omega\left(\frac{\log(T)}{\epsilon^2}\right).$$

- Hence, it remains to show that  $\mathbf{E}_{q_1}[n_{j,T}] = \Omega(\log(T)/\epsilon^2)$  for each  $j \notin J$ .



## Proof of Lemma 10 (contd.)

- $\mathbf{E}_{q_j}[T - n_{j,T}] = o((T^a)/\epsilon)$ .
  - Regret  $\epsilon$  is incurred for every comparison not involving  $b_j$ .
- By Markov's inequality,

$$q_j(\mathcal{E}_j) = q_j(\{T - n_{j,T} > T - \log(T)/\epsilon^2\}) \leq \frac{\mathbf{E}_{q_j}[T - n_{j,T}]}{T - \log(T)/\epsilon^2} = o(T^{a-1}).$$

- Choose a sufficiently large  $T$  so that  $q_j(\mathcal{E}_j) < 1/3$  for each  $j$ .

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For any event  $\mathcal{E}$  and distributions  $p, q$  with  $p(\mathcal{E}) \geq 1/3$  and  $q(\mathcal{E}) < 1/3$ ,

$$KL(p||q) \geq \frac{1}{3} \ln \left( \frac{1}{3q(\mathcal{E})} - \frac{1}{e} \right).$$

# Proof of Lemma 10 (contd.)

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For any event  $\mathcal{E}$  and distributions  $p, q$  with  $p(\mathcal{E}) \geq 1/3$  and  $q(\mathcal{E}) < 1/3$ ,

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- We have

$$KL(q_1||q_j) \geq \frac{1}{3} \ln \left( \frac{1}{o(T^{a-1})} \right) - \frac{1}{e} = \Omega(\log T).$$



# Proof of Lemma 10 (contd.)

- On the other hand, by the chain rule for KL-divergence,

$$KL(q_1||q_j) \leq \mathbf{E}_{q_1}[n_{j,T}] \cdot KL(1/2 + \epsilon||1/2 - \epsilon) \leq 16\epsilon^2 \cdot \mathbf{E}_{q_1}[n_{j,T}].$$

- If a comparison does not involve  $b_j$ , then the distribution on the comparison outcome will be the same under  $q_1$  and  $q_j$ .
- $KL(1/2 + \epsilon||1/2 - \epsilon)$ : the KL-divergence b/w two Bernoulli distributions  $\text{Ber}(1/2 + \epsilon), \text{Ber}(1/2 - \epsilon)$ .

## KL-divergence

For two probability mass functions  $p(x_1, \dots, x_r)$  and  $q(x_1, \dots, x_r)$ ,

$$KL(p(x_1, \dots, x_r)||q(x_1, \dots, x_r)) = \sum_{x_1} \dots \sum_{x_r} p(x_1, \dots, x_r) \log \frac{p(x_1, \dots, x_r)}{q(x_1, \dots, x_r)}.$$

- Hence,  $\mathbf{E}_{q_1}[n_{j,T}] = \Omega(\log(T)/\epsilon^2)$  for  $j \notin J$ .



